

# All of Statistics - Chapter 4 Solutions

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**1.**

Chebyshev's inequality gives  $\mathbb{P}(|X - \mu| \geq k\sigma) \leq 1/k^2$ . An exact calculation yields instead  $e^{-(1+k)}$ . To see this, note that  $\beta(\mu \pm k\sigma) = 1 \pm k$  and  $1 - k < 0$  so that

$$\mathbb{P}(|X - \mu| \leq k\sigma) = \mathbb{P}(X \leq \mu + k\sigma) = F(\mu + k\sigma) = 1 - e^{-(1+k)}$$

**2.**

$$\mathbb{P}(X \geq 2\lambda) = \mathbb{P}(X - \lambda \geq \lambda) = \mathbb{P}(|X - \lambda| \geq \lambda) \leq 1/\lambda.$$

**3.**

First, note that  $\mathbb{V}(\bar{X}) = \mathbb{V}(X_1)/n = p(1-p)/n$ . Chebyshev's inequality yields

$$\mathbb{P}(|\bar{X} - p| > \epsilon) \leq \frac{p(1-p)}{n\epsilon^2} \leq \frac{\max\{x(1-x) : 0 \leq x \leq 1\}}{n\epsilon^2} = \frac{1}{4n\epsilon^2}.$$

Next, note that

$$\mathbb{P}(|\bar{X} - p| \geq \epsilon) = \mathbb{P}(\bar{X} - p \geq \epsilon) + \mathbb{P}(\bar{X} - p \leq -\epsilon).$$

Let  $Y_i = (X_i - \mathbb{E}X_1)/n = (X_i - p)/n$  so that  $\bar{X} - p = \sum_i Y_i$ . Then,  $\mathbb{E}Y_i = 0$  and  $-p/n \leq Y_i \leq (1-p)/n$ . Hoeffding's inequality yields

$$\begin{aligned} \mathbb{P}(\bar{X} - p \geq \epsilon) &= \mathbb{P}\left(\sum_i Y_i \geq \epsilon\right) \leq \exp(-t\epsilon) \prod_i \exp\left(\frac{t^2}{8n^2}\right) \\ &= \exp\left(\frac{t^2}{8n} - t\epsilon\right) \leq \min_{t>0} \exp\left(\frac{t^2}{8n} - t\epsilon\right) = \exp(-2n\epsilon^2). \end{aligned}$$

Similarly,  $\mathbb{P}(\bar{X} - p \leq -\epsilon) = \mathbb{P}(\sum_i (-Y_i) \geq \epsilon) \leq \exp(-2n\epsilon^2)$ . It follows that

$$\mathbb{P}(|\bar{X} - p| \geq \epsilon) \leq 2 \exp(-2n\epsilon^2) = \frac{1}{1/2 + n\epsilon^2 + n^2\epsilon^4 + O(n^4)}$$

is tighter than the Chebyshev bound for sufficiently large  $n$ .

**4.**

**a)**

Applying our findings from Question 3,

$$\mathbb{P}(p \in C_n) = 1 - \mathbb{P}(p \notin C_n) \geq 1 - 2 \exp(-2n\epsilon_n^2) = 1 - 2 \exp\left(\log\left(\frac{\alpha}{2}\right)\right) = 1 - \alpha.$$

**b)**

TODO (Computer Experiment)

**c)**

The length of the interval is  $2\epsilon_n$ . This length is at most  $c > 0$  if and only if  $n \geq 2 \log(2/\alpha)/c^2$ .

TODO (Plot)

**5.**

As per the hint,

$$\begin{aligned} \mathbb{P}(|Z| > t) &= 2\mathbb{P}(Z \geq t) = \sqrt{\frac{2}{\pi}} \int_t^\infty \exp\left(-\frac{x^2}{2}\right) dx \\ &\leq \sqrt{\frac{2}{\pi}} \frac{1}{t} \int_t^\infty x \exp\left(-\frac{x^2}{2}\right) dx = \sqrt{\frac{2}{\pi}} \frac{1}{t} \exp\left(-\frac{t^2}{2}\right). \end{aligned}$$

**6.**

TODO (Plot)

**7.**

A linear combination of IID normal random variables is itself a normal random variable. Therefore,  $\bar{X}$  is a random variable with zero mean and variance  $1/n$ . Letting  $Z \sim N(0, 1)$ , Mill's inequality yields

$$\mathbb{P}(|\bar{X}| \geq t) = \mathbb{P}(|Z| \geq t\sqrt{n}) \leq \sqrt{\frac{2}{\pi}} \frac{1}{t\sqrt{n}} \exp\left(-\frac{t^2 n}{2}\right).$$

The above is tighter than the Chebyshev bound  $1/(t^2 n)$  for sufficiently large  $n$ .