Math 525: Lecture 3

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1 Random variables

Consider rolling two dice, corresponding to the sample space

$$\Omega = \{ (m, n) \colon 1 \le m, n \le 6 \}.$$

We can compute various numerical quantities based on the outcome of the rolls. For example, the sum of the two dice X(m,n) = m + n

or their product

Y(m,n) = mn.

As we will see shortly, both X and Y are examples of random variables on a probability space. At least intuitively, a random variable is any function that maps the outcome of an experiment (e.g., (m, n)) to a numerical value (e.g., m + n or mn).

Before we give a rigorous definition of a random variable, let's compute as a motivating example the probability that the two die sum to 5. Letting X(m, n) = m + n, this is simply

 $\mathbb{P}(\{\omega \in \Omega \colon X = 5\}) = \mathbb{P}(\{(1,4), (2,3), (3,2), (4,1)\}) = 4/36 = 1/9.$

To give the rigorous definition of a random variable, we review the concept of an inverse map.

Definition 1.1. Let $f: A \to B$ be a function. Let $H \subset B$. Let

$$f^{-1}(H) = \{a \in A \colon f(a) \in H\}.$$

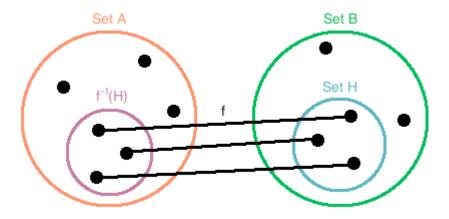
We call f^{-1} the *inverse map* of f. Note that the inverse map does not map points, but rather sets.

We are now ready to give the rigorous definition of a random variable. For the remainder, we will assume an underlying probability space $(\Omega, \mathcal{F}, \mathbb{P})$.

Definition 1.2. A random variable is a function $X: \Omega \to \mathbb{R}$ which satisfies, for all $B \in \mathcal{B}(\mathbb{R})$,

$$X^{-1}(B) \equiv \{\omega \in \Omega \colon X(\omega) \in B\} \in \mathcal{F}.$$

The Borel σ -algebra is rather large, so the above definition is not easy to check. The following proposition makes our lives a bit simpler.



The subset H of B has an inverse image f¹(H) which is a subset of A.

Figure 1: Inverse map

Proposition 1.3. Suppose \mathcal{G} generates the Borel σ -algebra (i.e., $\sigma(\mathcal{G}) = \mathcal{B}(\mathbb{R})$). Then, $X: \Omega \to \mathbb{R}$ is a random variable if and only if for each generating set $G \in \mathcal{G}$,

$$X^{-1}(G) \in \mathcal{F}.$$

Proof. We prove only the nontrivial direction. Let

$$\mathcal{M} = \left\{ B \subset \mathbb{R} \colon X^{-1}(B) \in \mathcal{F} \right\}.$$

By definition, we know that $\mathcal{G} \subset \mathcal{M}$. Therefore, $\sigma(\mathcal{G}) \subset \sigma(\mathcal{M})$. If \mathcal{M} is a σ -algebra, it follows that

$$\mathcal{B}(\mathbb{R}) = \sigma(\mathcal{G}) \subset \sigma(\mathcal{M}) = \mathcal{M},$$

as desired. To check that \mathcal{M} is a σ -algebra, verify the three properties:

1. $X^{-1}(\emptyset) = \emptyset \in \mathcal{F}.$ 2. If $B \in \mathcal{M}$, then $X^{-1}(B^c) = (X^{-1}(B))^c \in \mathcal{F}.$ 3. If $B_1, B_2, \ldots \in \mathcal{M}$, then $X^{-1}(\bigcup_{n \ge 1} B_n) = \bigcup_{n \ge 1} X^{-1}(B_n) \in \mathcal{F}.$

The above proposition is particularly useful when we take \mathcal{G} to be the set of intervals $\mathcal{G} = \{(-\infty, x] : x \in \mathbb{R}\}$. In this case...

Corollary 1.4. $X: \Omega \to \mathbb{R}$ is a random variable if and only if for each $x \in \mathbb{R}$,

$$\{\omega \in \Omega \colon X(\omega) \le x\} \in \mathcal{F}$$

Proof. If $G = (-\infty, x]$,

$$X^{-1}(G) = X^{-1}((-\infty, x]) = \{ \omega \in \Omega \colon X(\omega) \le x \} .$$

We will often use the above corollary to prove something is a random variable. To simplify notation, we often write

$$\{\omega \in \Omega \colon X(\omega) \le x\} = \{X \le x\}.$$

The above means that for a random variable X, $\mathbb{P}(\{X \leq x\})$ (i.e., the probability that X is at most x) is always well-defined. We define $\{X < x\}, \{X = x\}, \{X \geq x\}$, and $\{X > x\}$ similarly.

Proposition 1.5. Let X be a random variable. Then, $\{X < x\}, \{X = x\}, \{X \ge x\}, \{X > x\} \in \mathcal{F}$.

Proof. Let's just do the case of $\{X < x\}$. We can write

$$\{X < x\} = \bigcup_{\substack{q \in \mathbb{Q} \\ q > 0}} \{X \le x - q\},\$$

in which case we see that $\{X < x\}$ is nothing other than a countable union of sets of the form $\{X \le a\}$, which we know to be in \mathcal{F} .

Proposition 1.6. Let X and Y be random variables (on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$) and let a be a real number. Then, the following are also random variables:

- 1. aX.
- 2. X + Y.
- 3. XY.

4. Z defined by
$$Z(\omega) = \begin{cases} Y(\omega)/X(\omega) & \text{if } X(\omega) \neq 0\\ 0 & \text{if } X(\omega) = 0. \end{cases}$$

Proof. Let $x \in \mathbb{R}$ be arbitrary.

- 1. Note that $\{aX \le x\} = \{X \le x/a\}$. Since X is a random variable, $\{X \le x/a\} \in \mathcal{F}$.
- 2. It is sufficient to prove that $\{X + Y > x\} \in \mathcal{F}$ since $\{X + Y > x\} = \{X + Y \le x\}^c$. Note that

$$\{X + Y > x\} = \bigcup_{q \in \mathbb{Q}} \{X > q\} \cap \{Y > x - q\}.$$

In this form, it is clear that $\{X + Y > x\} \in \mathcal{F}$.

3. For a real number a, define

$$a^+ = \max\{a, 0\}$$
 and $a^- = \max\{-a, 0\}$

as its positive and negative parts. Since for any real number a we have $a = a^+ - a^-$, it follows that

$$XY = (X^{+} - X^{-})(Y^{+} - Y^{-}) = X^{+}Y^{+} + X^{+}Y^{-} - X^{-}Y^{+} + X^{-}Y^{-}.$$

Therefore, it is sufficient to consider the case in which X and Y are nonnegative. Moreover,

$$\{XY > x\} = \bigcup_{\substack{q \in \mathbb{Q} \\ q > 0}} \{X > q\} \cap \{Y > x/q\}.$$

4. It is sufficient to consider the case of Y = 1 since Y/X = Y(1/X). In this case,

$$\begin{split} \{Z \leq z\} &= \Omega \cap \{Z \leq z\} \\ &= (\{X \neq 0\} \cup \{X = 0\}) \cap \{Z \leq z\} \\ &= (\{X \neq 0\} \cap \{Z \leq z\}) \cup (\{X = 0\} \cap \{Z \leq z\}) \\ &= (\{X \neq 0\} \cap \{1/X \leq z\}) \cup (\{X = 0\} \cap \{0 \leq z\}) \,. \end{split}$$

First, note that $\{0 \le z\}$ is either equal to \emptyset or Ω , and hence $\{X = 0\} \cap \{0 \le z\} \in \mathcal{F}$. We leave verifying that $\{X \ne 0\} \cap \{1/X \le z\} \in \mathcal{F}$ as an exercise.

2 Distribution functions

Definition 2.1. The distribution function of a random variable X is the function $F \colon \mathbb{R} \to [0,1]$ defined by

$$F(x) = \mathbb{P}(\{X \le x\}).$$

The distribution function is sometimes called the *cumulative distribution function* (CDF).

Example 2.2. Flip a coin twice. Let X be the number of heads H which show up. Since

$$\mathbb{P}(\{TT\}) = 1/4$$
$$\mathbb{P}(\{HT, TH\}) = 2/4$$
$$\mathbb{P}(\{HH\}) = 1/4$$

we have that

$$F(x) = \begin{cases} 0 & \text{if } x < 0\\ 1/4 & \text{if } x < 1\\ 3/4 & \text{if } x < 2\\ 1 & \text{if } x \ge 2 \end{cases}$$

Remark 2.3. It now becomes clear why a random variable X requires $\{X \leq x\} \in \mathcal{F}!$ It is so that the dstribution function is well-defined at any point x.

In terms of the inverse function, note that

$$F(x) = \mathbb{P}(X^{-1}((-\infty, x]))$$

Proposition 2.4. Let X be a random variable and F be its distribution function. Then,

1. F is nondecreasing.

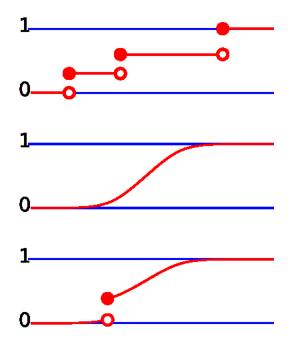


Figure 2: Examples of distribution functions

- 2. F is right continuous. That is, $F(x) = \lim_{y \downarrow x} F(y)$ for each x.
- 3. $\lim_{x\to\infty} F(x) = 0$ and $\lim_{x\to\infty} F(x) = 1$.

Proof. Let $A_x = \{X \leq x\}$ so that $F(x) = \mathbb{P}(A_x)$.

- 1. If $x \leq y$, then $A_x \subset A_y$ and hence $\mathbb{P}(A_x) \leq \mathbb{P}(A_y)$.
- 2. Let $(y_n)_n$ be a sequence converging to x from above. Then, we have $A_{y_1} \supset A_{y_2} \supset \cdots$ and hence $\lim_{n\to\infty} \mathbb{P}(A_{y_n}) = \mathbb{P}(\bigcap_{n\geq 1} A_{y_n}) = \mathbb{P}(A_x)$.
- 3. Note that $A_{-1} \supset A_{-2} \supset \cdots$ and hence $\lim_{n\to\infty} \mathbb{P}(A_{-n}) = \mathbb{P}(\bigcap_{n\geq 1}A_{-n}) = \mathbb{P}(\emptyset) = 0$. Similarly, $A_1 \subset A_2 \subset \cdots$ and hence $\lim_{n\to\infty} \mathbb{P}(A_n) = \mathbb{P}(\bigcup_{n\geq 1}A_n) = \mathbb{P}(\Omega) = 1$.

Since F is monotone, it follows that F has at most countably many discontinuities. We define

$$F(x-) = \lim_{y \uparrow x} F(x)$$

as the left-hand limit of F.

Proposition 2.5. Let X be a random variable with distribution function F. Then,

1. $\mathbb{P}(\{X < x\}) = F(x-).$ 2. $\mathbb{P}(\{X = x\}) = F(x) - F(x-).$ 3. If a < b, $\mathbb{P}(\{a < X \le b\}) = F(b) - F(a).$ 4. $\mathbb{P}(\{X > x\}) = 1 - F(x).$

The above implies that if F is continuous, then $\mathbb{P}(\{X = x\}) = 0$ for all x! This means that there is zero probability of any particular realization of a random variable. Positive probability is assigned only to ranges (e.g., $\{a < X \le b\}$).

3 Indicator random variables

Definition 3.1. Let $A \in \mathcal{F}$. The *indicator random variable on* A is the function

$$I_A(\omega) = \begin{cases} 1 & \text{if } \omega \in A \\ 0 & \text{if } \omega \notin A. \end{cases}$$

Note that for an indicator random variable I_A , we have $\mathbb{P}(\{I_A = 1\}) = \mathbb{P}(A)$ and $\mathbb{P}(\{I_A = 0\}) = \mathbb{P}(A^c) = 1 - \mathbb{P}(A)$.

Example 3.2. You play a game in which you roll a dice, and if the number you roll is greater than four, you get \$2. Otherwise, you lose \$1. How do we represent your winnings as a random variable?

Let $\Omega = \{1, \ldots, 6\}$. Since this is a finite sample space, we can safely take $\mathcal{F} = 2^{\Omega}$ and define \mathbb{P} by $\mathbb{P}(\{\omega\}) = 1/6$ (each outcome is equally as likely). The random variable corresponding to your winnings is

$$X(\omega) = 2I_{\{\omega>4\}}(\omega) - I_{\{\omega\le4\}}(\omega).$$

We'll often express this more succinctly as

$$X = 2I_{\{\omega > 4\}} - I_{\{\omega \le 4\}}.$$

Remark 3.3. There are many other common notations for indicator functions. It helps to be aware of these:

$$I_A(\omega) \equiv \mathbf{1}_A(\omega) \equiv \chi_A(\omega) \equiv [A](\omega).$$

4 Borel measurable functions

In closing this section, we introduce the notion of a Borel measurable function, which will give us our final piece of insight as to why we care about Borel σ -algebras.

Definition 4.1. We say $f : \mathbb{R} \to \mathbb{R}$ is a *Borel measurable function* (a.k.a. Borel function) if for each $B \in \mathcal{B}(\mathbb{R})$,

$$f^{-1}(B) \in \mathcal{B}(\mathbb{R}).$$

That is, the inverse image of a Borel set is also a Borel set.

Proposition 4.2. If $f : \mathbb{R} \to \mathbb{R}$ is continuous, it is Borel measurable.

Proof. Let

$$\mathcal{M} = \left\{ B \subset \mathbb{R} \colon f^{-1}(B) \in \mathcal{B}(\mathbb{R}) \right\}.$$

As usual, we can show that \mathcal{M} is a σ -algebra (check). Let \mathcal{G} be the set of open intervals in \mathbb{R} . Then, since f is continuous, $\mathcal{G} \subset \mathcal{M}$ and as such, $\mathcal{B}(\mathbb{R}) = \sigma(\mathcal{G}) \subset \sigma(\mathcal{M}) = \mathcal{M}$ as desired.

Exercise 4.3. Let X be a random variable and $f : \mathbb{R} \to \mathbb{R}$ be a Boreal measurable function. Show that Y defined by $Y(\omega) = f(X(\omega))$ is a random variable.

The above implies, most importantly, that taking the composition of a continuous function f and a random variable X yields a new random variable.